

# Kalman Filtering with Intermittent Observations: Tail Distribution and Critical Value

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**Abstract**—In this paper we analyze the performance of Kalman filtering for discrete-time linear Gaussian systems, where packets containing observations are dropped according to a Markov process modeling a Gilbert-Elliot channel. To address the challenges incurred by the loss of packets, we give a new definition of non-degeneracy, which is essentially stronger than the classical definition of observability, but much weaker than one-step observability, which is usually used in the study of Kalman filtering with intermittent observations. We show that the trace of the Kalman estimation error covariance under intermittent observations follows a power decay law. Moreover, we are able to compute the exact decay rate for non-degenerate systems. Finally we derive the critical value for non-degenerate systems based on the decay rate, improving upon the state of the art.

**Index Terms**—Estimation, Networked Control, Kalman filtering.

## I. INTRODUCTION

A large wealth of applications demands wireless communication among small embedded devices. Wireless Sensor Network (WSN) technology provides the architectural paradigm to monitor and control systems with a high degree of temporal and spatial granularity. Applications of sensor networks are becoming ubiquitous, ranging from environmental monitoring and control, to building automation, surveillance along with many others [1]. Given their low power nature and the requirement of long lasting deployment, communication between devices is limited in both range and reliability. Changes in the environment, such as the simple relocation of a large metal object in a room or the presence of people, inevitably affect the propagation properties of the wireless medium. Channels are time-varying, unreliable and difficult to characterize. Spurred by this consideration, our effort concentrates on the design and analysis of estimation and control algorithms over unreliable networks.

A substantial body of literature has been devoted to these issues in the past few years. In this paper, we want to revisit the problem of Kalman Filtering with intermittent observations, formulated by Sinopoli et al. [2]. In that paper, the authors analyze the problem of optimal state estimation for discrete-time linear Gaussian systems, under the assumption that observations are sent to the estimator via a memoryless analog erasure channel. This implies the existence of a non-unitary arrival probability associated with each packet. Consequently,

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some observations are inevitably lost. In this case, although the discrete Kalman Filter [3] is still the optimal estimator, the boundedness of its error depends on the arrival probabilities of the observation packets. In particular the authors prove the existence of a critical arrival probability  $p_c$ , below which the expectation of estimation error covariance matrix  $P_k$  of Kalman filter diverges. The authors are not able to compute the actual value of this critical probability for general linear systems, but provide upper and lower bounds, where the latter can be expressed in closed form as a function of the spectral radius of the dynamic matrix  $A$ . They are able to show that for special cases, for example when the observation matrix  $C$  is invertible, the upper bound coincides with the lower bound and hence the exact critical value is obtained.

Philosophically such a phenomenon can be seen as related to the well known uncertainty principle [4], [5], which states that the optimal long-range control of a linear system with uncertain parameters does not exist if the uncertainty exceeds a certain threshold. For Kalman filtering with intermittent observations, the uncertainty is incurred by the random packet loss and the optimal Kalman filter becomes unstable (i.e. the expectation of  $P_k$  is unbounded) if too many packets are dropped.

A significant amount of research effort has been made to analyze systems with intermittent observations. One interesting direction aims at characterizing the critical value for more general linear systems. Plarre and Bullo [6] relax the invertibility condition on the  $C$  matrix to  $C$  being invertible only on the observable subspace. In [7], the authors prove that if the eigenvalues of the system matrix  $A$  have distinct absolute values, then the lower bound is indeed the critical value. The authors also provide a counter example to show that in general the lower bound is not tight.

The drawback of the above approach is that the critical value only characterizes the boundedness of the expectation. To completely characterize the impact of lossy network on state estimation, it is much more desirable to calculate the probability distribution of estimation error covariance matrix  $P_k$  instead of only considering the boundedness of its expectation. In [8], the author gives a closed-form expression for cumulative distribution function of  $P_k$  when the system satisfies non-overlapping conditions. In [9], the authors provide a numerical method to calculate the eigen-distribution of  $P_k$  under the assumption that the observation matrix  $C$  is random and time varying. In [10], the authors consider the probability of  $P(P_k \leq X)$ <sup>1</sup>, and derive upper and lower bounds on such probability. However, only in specific cases these upper and

<sup>1</sup>Note that all the comparisons between matrices in this paper are in the sense of positive definite unless otherwise specified.

lower bounds coincide.

Other variations of the original problem are also considered. In [11], the authors introduce smart sensors, which send the local Kalman estimation instead of raw observation. In [12], a similar scenario is discussed where the sensor sends a linear combination of the current and previous measurements. A Markovian packet dropping model is introduced in [13] and a stability criterion is given. However, the authors can only find necessary and sufficient conditions for scalar systems. Only a sufficient condition is given for general systems. In [14], the authors study the case where the observation at each time splits into two parts, which are sent to the Kalman filter through two independent erasure channels. This work is further generalized by Garone et al. [15] to an arbitrary number of channels and extended to the control case. A much more general model, which considered packet drop, delay and quantization of measurements at the same time, is introduced by Xie and Shi [16].

In the meantime, significant efforts have been made to design estimation and control schemes over lossy networks, leveraging some of the results and methodologies mentioned above. Estimation of an unstable system is particularly important in control applications. Schenato et al. [17] show how an estimator rendered unstable by sensor data loss can in turn make the closed loop system unstable. Similar packet drop models have been successfully used in sensor selection problems for estimation in sensor networks. In [18], the authors consider a stochastic sensor scheduling scheme, which randomly selects one sensor to transmit observations at each time. In [19], the authors show how to design the packet arrival rate to balance the state estimation error and energy cost of packet transmission.

This paper breaks away from the existing body of literature and uses a novel approach to characterize the performance of Kalman filtering under packet losses. We study the tail distribution of  $\text{trace}(P_k)$ , i.e. how  $P(\text{trace}(P_k) > M)$  converges to 0 as  $M$  goes to infinity. The boundedness of the moments is then derived as a consequence of the tail distribution. This paper contains several significant contributions. Firstly, to the best of our knowledge, this is the first attempt to compute the decay rate of  $P(\text{trace}(P_k) > M)$ . This result has tremendous value of its own, as it provides the designer with a quantitative characterization on the quality of the estimator. Secondly, it derives the exact decay rate for a larger class of linear Gaussian systems, defined as non-degenerate. It illustrates the relationship between non-degeneracy and observability and argue that such condition, rather than the weaker notion of observability, is the appropriate one to check when the observation process is random. It is worth noticing that most literature [2], [6], [8], [10], [13] on Kalman filtering with intermittent observations requires the system to be one-step observable, which is stronger than non-degeneracy, in order to derive a tight condition on the boundedness of the filter. In this paper, for the case of Markovian packet loss model, we are able to relax this condition and provide tight boundedness conditions for all moments of the error covariance. As a special case, we are able to compute the critical value for i.i.d packet drops, showing that it coincides with the lower bound found

in [2] for non-degenerate systems. Lastly for non-degenerate systems we derive useful bounds for the Riccati equation, that can be used for any loss process. As a result, the proposed methodology can be applied to general packet loss models, provided that certain probabilities associated with the specific loss process can be computed.

The paper is organized in the following manner: Section II formulates the problem. Section III defines non-degenerate systems and compares this notion to observability and one-step observability. Section IV states several important inequalities on iterative Riccati and Lyapunov equations, which are used in Section V to derive the exact decay rate for non-degenerate systems. In Section VI we derive the critical value and boundedness conditions for higher moments of  $P_k$  based on the tail distribution, and compare our results with the existing literature. Finally Section VII concludes the paper.

## II. PROBLEM FORMULATION

Consider the following linear system

$$\begin{aligned} x_{k+1} &= Ax_k + w_k, \\ y_k &= Cx_k + v_k, \end{aligned} \quad (1)$$

where  $x_k \in \mathbb{C}^n$  is the state vector,  $y_k \in \mathbb{C}^m$  is the output vector,  $w_k \in \mathbb{C}^n$  and  $v_k \in \mathbb{C}^m$  are Gaussian random vectors<sup>2</sup> with zero mean and covariance matrices  $Q \geq 0$  and  $R > 0$ , respectively. Assume that the initial state  $x_0$  is also a Gaussian vector of mean  $\bar{x}_0$  and covariance matrix  $\Sigma \geq 0$ . Let  $w_i, v_i, x_0$  be mutually independent. Define  $|\lambda_1| \geq |\lambda_2| \geq \dots \geq |\lambda_n|$ , not necessarily distinct, as the eigenvalues of  $A$ .

Consider the case in which observations are sent to the estimator via a Gilbert-Elliot channel, where the packet arrival is modeled by a Markov process  $\{\gamma_k\}$ . According to this model, the measurement  $y_k$  sent at time  $k$  reaches its destination if  $\gamma_k = 1$ ; it is lost otherwise. Let  $\gamma_k$  be independent of  $w_k, v_k, x_0$ , i.e. the communication channel is independent of both process and measurement noises and let the transition matrix be

$$\begin{aligned} &\begin{bmatrix} P(\gamma_{k+1} = 0 | \gamma_k = 0) & P(\gamma_{k+1} = 1 | \gamma_k = 0) \\ P(\gamma_{k+1} = 0 | \gamma_k = 1) & P(\gamma_{k+1} = 1 | \gamma_k = 1) \end{bmatrix} \\ &= \begin{bmatrix} 1 - p_1 & p_1 \\ p_2 & 1 - p_2 \end{bmatrix}. \end{aligned}$$

Furthermore, we assume the Markov chain irreducible and stationary, i.e.  $0 < p_1 \leq 1$ ,  $0 < p_2 \leq 1$  and  $P(\gamma_0 = 0) = \dots = P(\gamma_k = 0) = p_2 / (p_1 + p_2)$ . This model also includes the i.i.d. packet drop model, as it can be seen as a special case of the Markovian model. It is in fact easy to prove if  $p_1 + p_2 = 1$ , then  $\gamma_k$ s are i.i.d. distributed and we define  $p \triangleq p_1$  in that case.

The Kalman Filter equations for this system were derived in [2] and take the following form:

$$\begin{aligned} \hat{x}_k &= \hat{x}_{k|k-1} + \gamma_k K_k (y_k - C \hat{x}_{k|k-1}), \\ P_k &= P_{k|k-1} - \gamma_k K_k C P_{k|k-1}, \end{aligned}$$

<sup>2</sup>Since we use  $A$  in its diagonal standard form, we assume every vector and matrix discussed in this paper to be in complex plane.

where

$$\begin{aligned}\hat{x}_{k+1|k} &= A\hat{x}_k, & P_{k+1|k} &= AP_kA^H + Q, \\ K_k &= P_{k|k-1}C^H(CP_{k|k-1}C^H + R)^{-1}, \\ \hat{x}_0 &= \bar{x}_0, & P_0 &= \Sigma,\end{aligned}$$

where  $(*)^H$  indicates the Hermitian operator. To simplify the notation, let us first define the functions

$$h(X, A, Q) \triangleq AXA^H + Q, \quad (2)$$

$$\tilde{g}(X, R) \triangleq X - XC^H(CXC^H + R)^{-1}CX, \quad (3)$$

$$g(X, A, Q, R) \triangleq \tilde{g}(h(X, A, Q), R). \quad (4)$$

When there is no confusion, we abbreviate  $h(X, A, Q)$  as  $h(X)$  and  $g(X, A, Q, R)$  as  $g(X)$ . It is easy to see that

$$P_{k+1} = \begin{cases} h(P_k), & \text{if } \gamma_{k+1} = 0 \\ g(P_k), & \text{if } \gamma_{k+1} = 1 \end{cases} \quad (5)$$

It can be proved that the following properties hold for both  $h$  and  $g$ :

- Proposition 1.** 1)  $h(X_1) \geq h(X_2)$  if  $X_1 \geq X_2$ .  
2)  $g(X_1) \geq g(X_2)$  if  $X_1 \geq X_2$ .  
3)  $h(X) \geq g(X)$ .

The proof can be found in Proposition 3 in the appendix. By definition,  $P_k$  is a random matrix since it is a function of  $\gamma_0, \dots, \gamma_k$ . In this paper we wish to analyze the tail behavior of the distribution of  $P_k$ , i.e., we want to know how likely it is for the Kalman filter to have a very large  $P_k$  due to packet loss. Let us first define

$$\varphi(M) \triangleq \sup_k P(\text{trace}(P_k) > M), \quad (6)$$

where  $M > 0$  is a scalar. Hence,  $\varphi(M)$  denotes the maximum probability of  $\text{trace}(P_k)$  to be larger than  $M$ . However, in general the exact value of  $\varphi(M)$  is hard to compute. In this paper we are more concerned with the decay rate under which  $\varphi(M)$  converges to 0 as  $M$  goes to infinity. Let us define the upper and lower decay rates respectively as

$$\bar{\phi} \triangleq \limsup_{M \rightarrow \infty} \frac{\log \varphi(M)}{\log M}, \quad \underline{\phi} \triangleq \liminf_{M \rightarrow \infty} \frac{\log \varphi(M)}{\log M}. \quad (7)$$

It is easy to see that since  $\varphi(M) \leq 1$ , both  $\bar{\phi}$  and  $\underline{\phi}$  are always non-positive. If  $\bar{\phi} = \underline{\phi}$ , then we define  $\phi \triangleq \bar{\phi} = \underline{\phi}$  as the decay rate.

**Remark 1.** If the decay rate  $\phi > -\infty$  is well defined, we can conclude from the definition that the following inequality holds for sufficient large  $M$ :

$$M^{\phi-\delta} \leq \varphi(M) \leq M^{\phi+\delta},$$

where  $\delta > 0$  can be arbitrarily small. As a result, we know the asymptotic behaviour of  $\varphi(M)$  when  $M$  approaches infinity. A slow decay rate would indicate that with high probability the filter has a large estimation error covariance  $P_k$  over time, while a fast decay rate indicates that such event is less likely.

In this paper we also want to characterize the conditions under which  $E(P_k^q)$  is uniformly bounded, where  $q \in \mathbb{N}$ . For the case where  $q = 1$  and  $\gamma_k$ s are independent, Sinopoli et

al. [2] prove the following existence theorem for the critical arrival probability:

**Theorem 1.** If  $(A, Q^{\frac{1}{2}})$  is controllable,  $(C, A)$  is detectable, and  $A$  is unstable, then there exists a critical value  $p_c \in [0, 1)$  such that <sup>3</sup>

$$\sup_k EP_k = +\infty \quad \text{for } 0 \leq p < p_c \quad \text{and} \quad \text{for some } P_0 \geq 0 \quad (8)$$

$$EP_k \leq M_{P_0} \quad \forall k \quad \text{for } p_c < p \leq 1 \quad \text{and} \quad \text{for all } P_0 \geq 0, \quad (9)$$

where  $M_{P_0} > 0$  depends on the initial condition  $P_0 \geq 0$ .

For simplicity, we say that  $EP_k$  is unbounded if  $\sup_k EP_k = +\infty$  or  $EP_k$  is bounded if there exists a uniform bound independent of  $k$ . In the following sections, we derive the decay rate of  $\varphi(M)$  and use it to characterize the boundedness of  $E(P_k^q)$ .

### III. NON-DEGENERATE SYSTEMS

In this section we introduce the concept of non-degeneracy and provide some insight on why the new definition is crucial in the analysis of Kalman filtering with intermittent observations. We also compare this new definition with the traditional definition of observability and one-step observability.

Before continuing on, we want to state the following assumptions on the system, which are assumed to hold throughout the rest of the paper:

(H1)  $(C, A)$  is detectable.

(H2)  $(A, Q^{1/2})$  is controllable.

(H3)  $A$  is unstable and diagonalizable<sup>4</sup>.

**Remark 2.** The first assumption is essential for the classical Kalman filter to have a bounded estimation error. The second assumption is used to guarantee the existence of the critical value, as per Theorem 1, and is usually satisfied by the majority of real systems. We further require  $A$  to be unstable because this is the only interesting case. If  $A$  is stable, one can prove that the Kalman filter still has a bounded estimation error even in the absence of observations. The study of estimation of unstable systems is particularly important for closed loop control, as an unstable estimator can render a closed loop system unstable. Finally the requirement of diagonalizability is used to define non-degenerate systems. While it is possible that such a requirement excludes some interesting systems, such as double integrators, most real systems are diagonalizable and therefore the results presented herein retain a great degree of generality. Although we believe that the results can be extended to Jordan forms, we expect the proofs to be lengthy, technical and not necessary insightful. Nonetheless, for the sake of completeness, we plan to remove this assumption and consider more general Jordan forms in future work.

Since we assume that  $A$  can be diagonalized, without loss of generality, we can assume it is already in diagonal form by performing a linear transformation on the system. Also since the eigenvalues of  $A$  can be complex, we use Hermitian instead of transpose in the rest of the paper.

<sup>3</sup>We use the notation  $\sup_k A_k = +\infty$  when there is no matrix  $M \geq 0$  such that  $A_k \leq M, \forall k$ .

<sup>4</sup>We allow complex eigenvalues.

We now are ready to define non-degenerate systems:

**Definition 1.** Consider the system  $(A, C)$  in its diagonal standard form, i.e.  $A = \text{diag}(\lambda_1, \dots, \lambda_n)$  and  $C = [C_1, \dots, C_n]$ . A block of the system is defined as a subsystem  $(A_{\mathcal{I}} = \text{diag}(\lambda_{i_1}, \dots, \lambda_{i_j}), C_{\mathcal{I}} = [C_{i_1}, \dots, C_{i_j}])$ ,  $1 \leq i_1 < \dots < i_j \leq n$ , where  $\mathcal{I} = \{i_1, \dots, i_j\} \subset \{1, \dots, n\}$  is the index set.

**Definition 2.** An equiblock is a block which satisfies  $\lambda_{i_1} = \dots = \lambda_{i_j}$ .

**Definition 3.** A quasi-equiblock is a block which satisfies  $|\lambda_{i_1}| = \dots = |\lambda_{i_j}|$ .

**Definition 4.** A system  $(A, C)$  is one-step observable if  $C$  is full column rank.

**Definition 5.** A diagonalizable system is non-degenerate if every quasi-equiblock of the system is one-step observable. It is degenerate if there exists at least one quasi-equiblock which is not one-step observable.

For example, if  $A = \text{diag}(2, -2)$  and  $C = [1, 1]$ , then the system is degenerate since it is a quasi-equiblock and it is not one-step observable. For  $A = \text{diag}(2, -2, 3, -3)$  and

$$C = \begin{bmatrix} 1 & 0 & 1 & 0 \\ 0 & 1 & 0 & 1 \end{bmatrix},$$

the two equiblocks are  $(\text{diag}(2, -2), I_2)$  and  $(\text{diag}(3, -3), I_2)$  and both of them are one-step observable. Thus, the system is non-degenerate.

To compare observability and non-degeneracy, we provide the following theorem:

**Theorem 2.** A system is observable if and only if every equiblock is one-step observable.

*Proof.* This is a direct result of Theorem 1 in [20] and the duality between observability and controllability.  $\square$

As a result, we can conclude that observability is weaker than non-degeneracy, since every equiblock is a quasi-equiblock. Therefore, a non-degenerate system is also observable. The converse is not true, as an observable system may be degenerate as shown in the example above.

On the other hand, non-degeneracy is obviously weaker than one-step observability. In fact, for a one-step observable system, the  $C$  matrix must have at least  $n$  independent rows, which implies  $y_k$  is at least a  $\mathbb{C}^n$  vector. While, for a non-degenerate system, the  $C$  matrix is required only to have the number of rows equal to the dimension of the largest quasi-equiblock. Enforcing one-step observability in large systems requires the use of a potentially high number of sensors and communications at each sampling period. Enforcing non-degeneracy in general requires the use of less sensors and communication bandwidth. We show that the same asymptotic performance is attained by both one-step observable and non-degenerate systems.

To summarize the above comparison, we have:

- 1) The system is observable if and only if every equiblock is one-step observable.

- 2) The system is non-degenerate if and only if every quasi-equiblock is one-step observable.
- 3) The system is one-step observable if  $C$  is full column rank.

Before proceeding, we wish to give some intuition on the importance of the concept of non-degeneracy for Kalman filtering with intermittent observations. A more rigorous discussion can be found in the next section. The main reason to introduce such concept is the loss of observability. It is well known that observability may be lost when discretizing a continuous time system or using a different sampling on a discrete time system, since different eigenvalues may rotate to the same point. The same thing happens when packet drops occur, which can be seen as a random sampling of the system. As a result, we need the stronger condition of non-degeneracy to ensure that the system remains always observable, regardless of the sampling rate. This is formalized in the following proposition:

**Proposition 2.** If the system is non-degenerate, then  $(A^q, C)$  is observable for any  $q \in \mathbb{R}$ .

*Proof.* We use Theorem 2. First consider an equiblock  $(A_{\mathcal{I}}^q, C_{\mathcal{I}})$ , where

$$A_{\mathcal{I}}^q = \text{diag}(\lambda_{i_1}^q, \dots, \lambda_{i_j}^q).$$

Since  $\lambda_{i_1}^q = \dots = \lambda_{i_j}^q$ , we know that

$$|\lambda_{i_1}| = \dots = |\lambda_{i_j}|.$$

As a result,  $(A_{\mathcal{I}}, C_{\mathcal{I}})$  is a quasi-equiblock for the original system. Hence, it is one-step observable, which implies that  $(A_{\mathcal{I}}^q, C_{\mathcal{I}})$  is also one-step observable. Therefore, all the equiblocks of  $(A^q, C)$  are one-step observable. Using Theorem 2 we conclude the proof.  $\square$

In the rest of the paper, we compute the exact decay rate for non-degenerate systems, given by

$$\phi = \frac{\log(1 - p_1)}{2 \log |\lambda_1|}. \quad (10)$$

#### IV. GENERAL INEQUALITIES ON ITERATIVE RICCATI AND LYAPUNOV EQUATIONS

Before deriving the decay rate, we establish several inequalities on the iterative Riccati and Lyapunov equations, which are used in the next section. We want to emphasize that such inequalities are derived independently of the packet loss model and can therefore be used in a more general context.

To simplify notation, let us define

$$\begin{aligned} g^1(X) &= g(X), & g^i(X) &= g(g^{i-1}(X)), \\ h^1(X) &= h(X), & h^i(X) &= h(h^{i-1}(X)). \end{aligned}$$

where  $g, h$  are defined in (4). Moreover, we simply write the composition of  $g^i$  and  $h^j$  as

$$g^i h^j(X) \triangleq g^i(h^j(X)).$$

The first inequality provides a lower bound for  $h^i(X)$ :

**Theorem 3.** Suppose the system satisfies assumptions (H1)–(H3), then the following inequality holds

$$\text{trace}(h^i(X)) \geq \underline{\alpha} |\lambda_1|^{2i}, \quad i \geq n, \quad (11)$$

where  $X$  is positive semidefinite,  $\underline{\alpha} > 0$  is a constant independent of  $X$  and  $i$  and  $|\lambda_1|$  is the spectral radius of  $A$ .

*Proof.* First let us consider  $h^n(X)$ . From the definition of  $h$ , we can simply write it as

$$h^n(X) = Q + AQA^H + \dots + A^{n-1}QA^{(n-1)H} + A^nXA^{nH},$$

where  $A^{nH} = (A^H)^n$ . Since we already assumed that  $(A, Q^{1/2})$  is controllable, we know that  $Q + \dots + A^{n-1}QA^{(n-1)H}$  is strictly positive definite. Hence, it is possible to find  $\underline{\alpha} > 0$ , such that

$$Q + AQA^H + \dots + A^{n-1}QA^{(n-1)H} > \underline{\alpha} |\lambda_1|^{2n} I_n > 0.$$

Now consider  $h^i(X)$ , where  $i \geq n$ ,

$$\begin{aligned} h^i(X) &= Q + AQA^H + \dots + A^{i-1}QA^{(i-1)H} + A^iXA^{iH} \\ &\geq A^{i-n}QA^{(i-n)H} + \dots + A^{i-1}QA^{(i-1)H} \\ &= A^{i-n} \left( \sum_{j=0}^{n-1} A^jQA^{jH} \right) A^{(i-n)H} \\ &\geq \underline{\alpha} |\lambda_1|^{2n} A^{i-n} A^{(i-n)H}. \end{aligned}$$

Taking the trace on both sides and using the fact that  $A$  is a diagonal matrix, we have

$$\text{trace}(h^i(X)) \geq \underline{\alpha} |\lambda_1|^{2n} \left( \sum_{j=1}^n |\lambda_j|^{2i-2n} \right) \geq \underline{\alpha} |\lambda_1|^{2i},$$

which concludes the proof.  $\square$

The following theorem characterizes the upper bound:

**Theorem 4.** Consider a system that satisfies assumptions (H1)–(H3). If the unstable part is non-degenerate, then the following inequality holds:

$$\text{trace}(h^{i_1}gh^{i_2-1}g \dots h^{i_l-1}gh^{i_{l+1}-1}(\Sigma)) \leq \bar{\alpha} \prod_{j=1}^l (|\lambda_j| + \varepsilon)^{2i_j}, \quad (12)$$

where,  $|\lambda_1| \geq \dots \geq |\lambda_l|$  are all the unstable eigenvalues of the system,  $\bar{\alpha} > 0$  is a constant independent of  $i_j$  and  $\varepsilon > 0$  can be arbitrarily small.

*Proof.* The proof is quite long and is reported in the appendix.  $\square$

**Remark 3.** Suppose that  $l$  observations are received at times  $k_1 > k_2 > \dots > k_l$ , where  $k \geq k_1$  and  $k_l \geq 0$ . To simplify notations, let us define  $k_0 = k$  and  $k_{l+1} = 0$  and  $i_j = k_{j-1} - k_j$ ,  $j = 1, \dots, l+1$ . From the definition of  $h$  and  $g$ , we know that

$$P_k = h^{i_1}gh^{i_2-1}g \dots h^{i_l-1}g(P_{k_{l-1}}),$$

and

$$P_{k_{l-1}} \leq h^{i_{l+1}-1}(\Sigma).$$

From Theorem 4, we know that the trace of  $P_k$  is bounded by  $\bar{\alpha} \prod_{j=1}^l (|\lambda_j| + \varepsilon)^{2i_j}$ . In other words, no matter how many packets are lost from time 1 to time  $k_l - 1$ , we can always have a bounded estimation error as long as  $l$  packets arrive between time  $k_l$  to time  $k$ . In the classical setting of perfect communications, it is easy to derive from the observability Grammian that the estimation error is bounded when  $n$  sequential packets are received, provided that the system is detectable. Hence, our result can be seen as a generalization of the classical result, since we do not require sequential packets arrivals. However, we do need to enforce the requirement of non-degeneracy, instead of observability, on the system.

To see why the requirement of non-degeneracy is crucial for (12) to hold, let us consider the system, where

$$A = \begin{bmatrix} 0 & \sqrt{2} \\ \sqrt{2} & 0 \end{bmatrix}, C = [1 \quad 0], Q = \Sigma = I_2, R = 1.$$

It is easy to check that the system satisfies assumptions (H1)–(H3). However it is degenerate. Let  $X = \text{diag}(a, b)$ . Simple algebraic calculation yields

$$h(X) = \text{diag}(2b + 1, 2a + 1),$$

$$g(X) = \text{diag}((2b + 1)/(2b + 2), 2a + 1) \geq \text{diag}(0, 2a + 1)$$

As a result, it is easy to check that

$$h^{i_1}gh^{i_2-1}gh^{i_3-1}(\Sigma) \geq \text{diag}(0, 2^{i_1+i_2+i_3+1} - 1)$$

when both  $i_1$  and  $i_2$  are even. Hence inequality (12) does not hold since the right hand side depends not only on  $i_1$  and  $i_2$  but also  $i_3$ . In other word, we cannot bound the estimation error by a function of  $i_1$  and  $i_2$  even when two packets have arrived.

Also note that in the statement of the theorem we do not assume any channel failure models. As a result, the above theorem can also be used to analyze the performance of Kalman filtering with other packet drop models.

## V. DECAY RATE FOR NON-DEGENERATE SYSTEMS

In this section, we use Theorem 3 and 4 to derive the decay rate for non-degenerate systems. Let us first define the following event:

$$E_{k,a}^N = \{\text{No more than } a \text{ packets arrive between time } k \text{ and time } k - N + 1\},$$

where  $k \geq N$ .

**Theorem 5.** Suppose the system satisfies assumptions (H1)–(H3), then

$$\underline{\phi} \geq \frac{\log(1 - p_1)}{2 \log |\lambda_1|}. \quad (13)$$

*Proof.* If  $|\lambda_1| = 1$ , then (13) becomes

$$\underline{\phi} \geq -\infty,$$

which always holds. Hence, without loss of generality, we assume  $|\lambda_1| > 1$ . Set  $M$  large enough and choose  $N \in \mathbb{N}$  such that

$$\underline{\alpha} |\lambda_1|^{2N-2} \leq M < \underline{\alpha} |\lambda_1|^{2N}.$$

Since  $|\lambda_1| > 1$ , we can always find such  $N$ . Let us consider the estimation error at time  $k > N$ . Suppose the observations made at time  $k - N + 1, \dots, k$  are all lost, then by definition of  $h$ , we know that

$$P_k = h^N(P_{k-N}).$$

From Theorem 3 we know that

$$\text{trace}(P_k) \geq \underline{\alpha}|\lambda_1|^{2N} > M.$$

As a result, when packets  $k - N + 1$  to  $k$  are all lost,  $\text{trace}(P_k) > M$ . Hence

$$\varphi(M) = \sup_k P(\text{trace}(P_k) > M) \geq P(E_{k,0}^N)$$

The probability on the right hand side can be easily computed from the Markovian packet loss assumption:

$$\begin{aligned} P(E_{k,0}^N) &= P(\gamma_{k-N+1} = \dots = \gamma_k = 0) \\ &= P(\gamma_{k-N+1} = 0) \prod_{i=k-N+1}^{k-1} P(\gamma_{i+1} = 0 | \gamma_i = 0) \\ &= \frac{p_2}{p_1 + p_2} (1 - p_1)^{N-1}. \end{aligned}$$

Since we already assumed that

$$\underline{\alpha}|\lambda_1|^{2N-2} \leq M,$$

we know that

$$N \leq \frac{\log M - \log \underline{\alpha}}{2 \log |\lambda_1|} + 1.$$

As a result,

$$\begin{aligned} \underline{\phi} &= \liminf_{M \rightarrow \infty} \frac{\log \varphi(M)}{\log M} \\ &\geq \liminf_{M \rightarrow \infty} \frac{\log(1 - p_1) \log M - \log \underline{\alpha}}{\log M} \frac{\log M - \log \underline{\alpha}}{2 \log |\lambda_1|} \\ &\quad + \liminf_{M \rightarrow \infty} \frac{\log p_2 - \log(p_1 + p_2)}{\log M} \\ &= \frac{\log(1 - p_1)}{2 \log |\lambda_1|}, \end{aligned}$$

which concludes the proof.  $\square$

**Remark 4.** The above theorem indicates that the distribution of  $\text{trace}(P_k)$  follows a power decay rule as it can decay at most  $M^{\underline{\phi} - \delta}$  fast when  $A$  is strictly unstable. Also since we do not assume non-degeneracy in the proof, the result is valid for general linear systems satisfying assumptions (H1) – (H3).

Now let us consider an upper bound for  $\bar{\phi}$ , which is given by the following theorem:

**Theorem 6.** Consider a system which satisfies assumptions (H1) – (H3) and the unstable part is non-degenerate, then the following inequality holds:

$$\bar{\phi} \leq \frac{\log(1 - p_1)}{2 \log |\lambda_1|}. \quad (14)$$

*Proof.* First we claim that

$$\text{trace}(P_k) \leq \beta |\lambda_1|^{2k},$$

where  $\beta > 0$  is a constant independent of  $k$ . It is clear that the worst case happens when all the packets from time 1 to time  $k$  are lost. Hence

$$P_k \leq h^k(\Sigma) = \sum_{j=0}^{k-1} A^j Q A^{jH} + A^k \Sigma A^{kH}.$$

Now we can find a constant  $\beta' > 0$ , such that  $Q \leq \beta' I_n$  and  $\Sigma \leq \beta' I_n$ . Therefore, we have

$$\begin{aligned} \text{trace}(P_k) &\leq \beta' \text{trace}\left(\sum_{j=0}^{k-1} A^j A^{jH}\right) = \beta' \sum_{i=1}^n \frac{|\lambda_i|^{2k+2} - 1}{|\lambda_i|^2 - 1} \\ &\leq n \beta' \frac{|\lambda_1|^{2k+2} - 1}{|\lambda_1|^2 - 1} \leq \frac{n \beta' |\lambda_1|^2}{|\lambda_1|^2 - 1} |\lambda_1|^{2k}, \end{aligned}$$

where we use the fact that  $A$  is diagonal. Hence,  $\text{trace}(P_k) \leq \beta |\lambda_1|^{2k}$ , where  $\beta = (n \beta' |\lambda_1|^2) / (|\lambda_1|^2 - 1)$ .

Now let us fix  $\varepsilon > 0$  defined in Theorem 4 and  $M$  large enough. Without loss of generality, we also assume  $\bar{\alpha} > \beta$ , where  $\bar{\alpha}$  is also defined in Theorem 4, since we can always pick a sufficiently large  $\bar{\alpha}$ . Choose  $N$  such that

$$\bar{\alpha}(|\lambda_1| + \varepsilon)^{2N} \leq M < \bar{\alpha}(|\lambda_1| + \varepsilon)^{2N+2}. \quad (15)$$

Since we assumed that  $\bar{\alpha} > \beta$ , we know that if  $k \leq N$ , then

$$\text{trace}(P_k) \leq \beta |\lambda_1|^{2k} \leq \beta |\lambda_1|^{2N} < M.$$

As a result,  $P(\text{trace}(P_k) > M) = 0$ , when  $k \leq N$ . Therefore

$$\varphi(M) = \sup_{k > N} P(\text{trace}(P_k) > M).$$

Now let us consider the case where  $k > N$ . Suppose we have received more than  $l$  packets from time  $k - N + 1$  to time  $k$ . Assume that the latest packet is received at time  $k_1$ , and the second latest packet is received at time  $k_2$ , and so on. To simplify notation, define  $k_0 = k$ ,  $k_{l+1} = 0$  and  $i_j = k_{j-1} - k_j$ ,  $j = 1, \dots, l+1$ . From the definition of  $h$  and  $g$ , we know that

$$P_k = h^{i_1} g h^{i_2-1} g \dots h^{i_l-1} g(P_{k_{l-1}}),$$

and

$$P_{k_{l-1}} \leq h^{i_{l+1}-1}(\Sigma).$$

From Theorem 4, we know that

$$\begin{aligned} \text{trace}(P_k) &= \text{trace}(h^{i_1} g h^{i_2-1} g \dots h^{i_l-1} g h^{i_{l+1}-1}(\Sigma)) \\ &\leq \bar{\alpha} \prod_{j=1}^l (|\lambda_j| + \varepsilon)^{2i_j} \leq \bar{\alpha} (|\lambda_1| + \varepsilon)^{2(k-k_l)}. \end{aligned}$$

Using the fact that more than  $l$  packets are received from time  $k - N + 1$  to time  $k$ , we know that

$$k_l \geq k - N + 1 \Rightarrow k - k_l < N,$$

which implies that

$$\text{trace}(P_k) < \bar{\alpha} (|\lambda_1| + \varepsilon)^{2N} \leq M.$$

As a result, if more than  $l$  packets are received from time  $k - N + 1$  to time  $k$ , then  $\text{trace}(P_k) \leq M$ . Therefore

$$P(\text{trace}(P_k) > M) \leq P(E_{k,l-1}^N).$$

Now let us estimate the probability on the right hand side. Unlike the  $P(E_{k,0}^N)$ , the exact value of such probability cannot be easily computed. As a result we only focus on the upper bound of  $P(E_{k,l-1}^N)$ . We know that there are totally  $2^N$  possible combinations of  $\gamma_{k-N+1}, \dots, \gamma_k$  and  $\sum_{i=0}^{l-1} \binom{N}{i}$  of them belong to the event  $E_{k,l-1}^N$ . Suppose that  $\gamma'_{k-N+1}, \dots, \gamma'_k$  is one of the eligible combinations. As a result

$$P(\gamma'_{k-N+1}, \dots, \gamma'_k) = P(\gamma'_{k-N+1}) \prod_{i=k-N}^k P(\gamma'_i | \gamma'_{i-1}).$$

Since there are at most  $l-1$   $\gamma'_i$ s which are 1, one can prove there exist at least  $N-2l+1$  pairs of  $(\gamma'_i, \gamma'_{i+1})$ , which are both 0. As a result, the probability to get  $(\gamma'_{k-N+1}, \dots, \gamma'_k)$  is upper bounded by

$$P(\gamma'_{k-N+1}, \dots, \gamma'_k) \leq (1-p_1)^{N-2l+1}.$$

Therefore

$$\begin{aligned} P(E_{k,l-1}^N) &\leq \sum_{i=0}^{l-1} \binom{N}{i} (1-p_1)^{-2l+1} (1-p_1)^N \\ &= \text{Poly}(N) (1-p_1)^N, \end{aligned}$$

where

$$\text{Poly}(N) = \sum_{i=0}^{l-1} \binom{N}{i} (1-p_1)^{-2l+1}$$

is a  $(l-1)^{\text{th}}$  polynomial of  $N$ . Therefore,

$$\begin{aligned} \bar{\phi} &= \limsup_{M \rightarrow \infty} \frac{\log \varphi(M)}{\log M} \\ &\leq \limsup_{M \rightarrow \infty} \frac{\log[\text{Poly}(N)]}{\log M} + \limsup_{M \rightarrow \infty} \frac{N \log(1-p_1)}{\log M}. \end{aligned}$$

For the first term, from the first inequality in (15), we have

$$\limsup_{M \rightarrow \infty} \left| \frac{\log[\text{Poly}(N)]}{\log M} \right| \leq \limsup_{N \rightarrow \infty} \left| \frac{\log[\text{Poly}(N)]}{2N \log(|\lambda_1| + \varepsilon)} \right| = 0.$$

For the second term, using the second inequality in (15), it is easy to establish the following inequality

$$N > \frac{\log M}{2 \log(|\lambda_1| + \varepsilon)} - 1.$$

Therefore,

$$\begin{aligned} &\limsup_{M \rightarrow \infty} \frac{N \log(1-p_1)}{\log M} \\ &\leq \limsup_{M \rightarrow \infty} \left( \frac{\log M}{2 \log(|\lambda_1| + \varepsilon)} - 1 \right) \frac{\log(1-p_1)}{\log M} \\ &= \frac{\log(1-p_1)}{2 \log(|\lambda_1| + \varepsilon)}. \end{aligned}$$

Thus, we can conclude that

$$\bar{\phi} \leq \frac{\log(1-p_1)}{2 \log(|\lambda_1| + \varepsilon)}.$$

If we let  $\varepsilon$  go to 0, we can conclude the proof.  $\square$

Combining the above two Theorems, we have

**Theorem 7.** Consider a system which satisfies assumptions (H1) – (H3) and assume the unstable part non-degenerate, then the decay rate is given by the following equality:

$$\phi = \frac{\log(1-p_1)}{2 \log |\lambda_1|}. \quad (16)$$

**Remark 5.** The results we derived here can be used for more general packet drop models, as long as  $P(E_{k,0}^n)$  and  $P(E_{k,l-1}^n)$  can be computed.

**Remark 6.** For a non-degenerate system, the decay rate is a decreasing function of  $p_1$ , which is very intuitive since larger  $p_1$  indicates the channel is more capable of recovering from a bad state. However this may not be true for a degenerate system. For example, consider the system  $A = \text{diag}(2, -2)$ ,  $C = [1, 1]$  and transition matrix

$$\begin{aligned} &\begin{bmatrix} P(\gamma_{k+1} = 0 | \gamma_k = 0) & P(\gamma_{k+1} = 1 | \gamma_k = 0) \\ P(\gamma_{k+1} = 0 | \gamma_k = 1) & P(\gamma_{k+1} = 1 | \gamma_k = 1) \end{bmatrix} \\ &= \begin{bmatrix} 1-p_1 & p_1 \\ 1 & 0 \end{bmatrix}. \end{aligned}$$

If  $p_1 = 1$ , then the sequence of  $\gamma_k$ s becomes 101010... Hence, it is equivalent to sampling the system only at even (or odd depends on the  $\gamma_0$ ) time and it is easy to see that such a system is unobservable. As a result  $\varphi(M) = 1$  for any  $M$  and the decay rate is 0. However, if  $p_1$  is not 1, then we can get several consecutive 0s in the sequence of  $\gamma_k$ . As a result, it is possible to get observations made both at odd and even time (for example it is possible to get "1001" for  $\gamma_0$  to  $\gamma_3$ ) and attain  $\phi \leq 0$ .

As discussed in the Section III, we know that the following inequality holds when  $M$  is sufficient large:

$$M^{\phi-\delta} \leq \sup_k P(\text{trace}(P_k) > M) \leq M^{\phi+\delta},$$

where  $\delta > 0$  can be arbitrarily small. Such result is very useful to characterize the tail distribution of  $P_k$ . In the next section, we apply this result to derive the boundedness conditions for higher moments of  $P_k$  and the critical value for the i.i.d. packet drop model.

## VI. CRITICAL VALUE AND BOUNDEDNESS OF HIGHER MOMENTS

In this section we want to derive the critical value and boundedness conditions of higher moments for Kalman filtering with intermittent observations based on the decay rate we derive in Section V, which is given by the following theorem:

**Theorem 8.** Consider a system which satisfies assumptions (H1) – (H3) and assume the unstable part non-degenerate. Let  $q \in \mathbb{N}$ .  $E[P_k^q]$  is uniformly bounded if

$$p_1 > 1 - \frac{1}{|\lambda_1|^{2q}}. \quad (17)$$

It is unbounded for some initial condition  $\Sigma$  if

$$p_1 < 1 - \frac{1}{|\lambda_1|^{2q}}. \quad (18)$$

*Proof.* Let us first prove the case where  $q = 1$ . It is easy to see that  $EP_k$  is uniformly bounded if and only if  $E[\text{trace}(P_k)]$  is uniformly bounded. As a result, we only focus on the boundedness of  $E[\text{trace}(P_k)]$ . Now let us first prove that if  $p_1 > 1 - |\lambda_1|^{-2}$ , then  $E[\text{trace}(P_k)]$  is uniformly bounded. Since  $\text{trace}(P_k)$  is non-negative, it is easy to see that the following equality holds

$$E[\text{trace}(P_k)] = \int_0^\infty P(\text{trace}(P_k) > M) dM.$$

Taking the supremum on both sides yields

$$\begin{aligned} \sup_k E[\text{trace}(P_k)] &= \sup_k \int_0^\infty P(\text{trace}(P_k) > M) dM \\ &\leq \int_0^\infty \varphi(M) dM. \end{aligned}$$

Now by the definition of decay rate  $\phi$ , we know that

$$\varphi(M) \leq M^{\phi+\delta},$$

provided that  $M$  is sufficient large and  $\delta > 0$  can be arbitrarily small. Hence, it is easy to check that the following condition is sufficient for  $\text{trace}(EP_k)$  to be uniformly bounded:

$$\phi + \delta < -1,$$

which is equivalent to

$$\frac{\log(1 - p_1)}{2 \log |\lambda_1|} < -1 - \delta.$$

Manipulating the above equations we know the following condition is sufficient

$$p_1 > 1 - |\lambda_1|^{-2-2\delta}.$$

Since  $\delta$  can be arbitrarily close to 0, we can conclude that

$$p_1 > 1 - |\lambda_1|^{-2},$$

is sufficient for  $E(P_k)$  to be uniformly bounded. Then let us prove that  $p_1 \leq 1 - |\lambda_1|^{-2}$  is sufficient for an unbounded  $E[\text{trace}(P_k)]$ . Since  $\text{trace}(P_k)$  is non-negative, it is easy to see that

$$E[\text{trace}(P_k)] \geq M \times P(\text{trace}(P_k) \geq M), \text{ for all } M.$$

Take the supremum over  $k$  and  $M$  on both side we have

$$\sup_k E[\text{trace}(P_k)] \geq \sup_M [M \times \varphi(M)]. \quad (19)$$

Since we know that

$$\varphi(M) \geq M^{\phi-\delta},$$

when  $M$  is sufficient large, the right hand side of (19) is unbounded if

$$1 + \phi - \delta > 0,$$

which is equivalent to

$$p_1 < 1 - |\lambda_1|^{-2+2\delta}.$$

Since  $\delta$  can be arbitrarily small, we can conclude that

$$p_1 < 1 - |\lambda_1|^{-2},$$

is sufficient for an unbounded  $E(P_k)$ .

Now let us consider arbitrary  $q$ . Suppose the eigenvalues of  $P_k$  are  $\xi_1 \geq \xi_2 \geq \dots \geq \xi_n$ . Hence,

$$\sum_{i=1}^n \xi_i = \text{trace}(P_k), \quad \sum_{i=1}^n \xi_i^q = \text{trace}(P_k^q).$$

Since  $P_k$  is positive semidefinite,  $\xi_i$  are non-negative. As a result, it is easy to prove that

$$\text{trace}(P_k^q) \leq \text{trace}(P_k)^q, \quad (20)$$

and the equality holds only when  $\xi_1 = \text{trace}(P_k)$  and  $\xi_2 = \dots = \xi_n = 0$ . Moreover,

$$\text{trace}(P_k^q) \geq n \left( \frac{\text{trace}(P_k)}{n} \right)^q = n^{1-q} \times \text{trace}(P_k)^q, \quad (21)$$

and the equality holds only when  $\xi_1 = \dots = \xi_n = \text{trace}(P_k)/n$ . Let us define

$$\varphi_q(M) \triangleq \sup_k P(\text{trace}(P_k^q) > M),$$

and similarly

$$\bar{\phi}_q \triangleq \limsup_{M \rightarrow \infty} \frac{\log \varphi_q(M)}{\log M}, \quad \underline{\phi}_q \triangleq \liminf_{M \rightarrow \infty} \frac{\log \varphi_q(M)}{\log M}.$$

Moreover  $\phi_q \triangleq \bar{\phi}_q$  if  $\bar{\phi}_q = \underline{\phi}_q$ . It is easy to see from (20) and (21) that

$$\phi_q = \phi/q.$$

Therefore, following the same argument used for the case  $q = 1$ , we conclude the proof.  $\square$

Since independent packet drop is a special case for the Markovian packet drop, we have the following corollary:

**Corollary 1.** Consider a system which satisfies assumptions (H1)–(H3) and the unstable stable part is non-degenerate. If  $\gamma_k$ s are i.i.d. distributed and  $P(\gamma_k = 1) = p$ , then the critical value of the system is given by

$$p_c = 1 - |\lambda_1|^{-2} \quad (22)$$

Before finishing this section, we want to compare our result on critical value, with the results from [2] and [6]. Let us assume that  $A = \text{diag}(\lambda_1, \dots, \lambda_n)$  and  $C = [C_1, \dots, C_n]$ . In [2], the authors derived the exact critical value in two cases:

- 1)  $C$  is invertible.
- 2)  $A$  has only one unstable eigenvalue.

If  $C$  is invertible, then the system is one-step observable and thus non-degenerate. For the second case, suppose that  $\lambda_1$  is the only unstable eigenvalue, it is easy to see that  $C_1$  could not be a zero vector otherwise the system would not be detectable. As a result, the unstable part, which is defined as the block  $A_{\mathcal{I}} = \lambda_1$ ,  $C_{\mathcal{I}} = C_1$ , is non-degenerate by definition.

In [6], the authors derive the critical value under the assumption that  $C$  is invertible on the observable subspace. Since the system is detectable, all the unstable eigenvalues must be observable. Suppose that  $\lambda_1, \dots, \lambda_l$  are unstable eigenvalues and  $\lambda_{l+1}, \dots, \lambda_{l'}$  are stable and observable eigenvalues. Then matrix  $[C_1, \dots, C_{l'}]$  must be full column rank



since  $C$  is invertible on the observable space, which implies that  $[C_1, \dots, C_l]$  must be also full column rank. Hence, the unstable part of the system is one-step observable and thus non-degenerate.

In conclusion, for diagonalizable systems, all the cases discussed in [2] and [6] are included in Corollary 1.

**Remark 7.** Note that the critical value for degenerate systems is not  $1 - |\lambda_1|^{-2}$  in general, as in shown in [7]. This is caused by the loss of observability incurred by packet drops. As a result, some packets contain only redundant information which does not improve the estimation on some mode of the system. In this case more packets may be needed to obtain bounded estimation.

## VII. CONCLUSIONS AND FUTURE WORK

In this paper we address the problem of state estimation for a discrete-time linear Gaussian system where observations are communicated to the estimator via an erasure channel. We were able to characterize the tail distribution of the estimation error covariance for Markovian packet losses and compute its decay rate. We then used this to compute the value of the critical probability for each of its moments. We introduced the concept of non-degeneracy and show how such concept, rather than the weaker notion of observability, is the appropriate property to check when observations are sampled according to a stochastic process. Finally, as a special case, we were able to derive the critical value for the case of i.i.d. packet losses, showing that it is indeed the lower bound found in [2].

This analysis does not address degenerate and non diagonalizable systems. Although some important systems fall within this class, we argue that our analysis covers most systems and can be extended, with considerable effort, to systems that admit non diagonal Jordan forms. As for degenerate systems, some preliminary results are contained in [7]. We plan to complete the characterization of the problem by generalizing our approach in both directions.

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## IX. APPENDIX

This section is devoted to proving Theorem 4. Before continuing on, we want to state the following properties of functions  $g$  and  $h$ , which are used in the proof:

**Proposition 3.** Let

$$\begin{aligned} h(X, A, Q) &= AXA^H + Q, \\ \tilde{g}(X, R) &= X - XC^H(CXC^H + R)^{-1}CX, \\ g(X, A, Q, R) &= \tilde{g}(h(X, A, Q), R), \end{aligned}$$

where  $X$ ,  $Q$  and  $R$  are defined in Section II. Then the following propositions hold:

- 1)  $h$  is an increasing function of  $X$  and  $Q$ <sup>5</sup>.
- 2)  $g$  is an increasing function of  $X$ ,  $Q$ ,  $R$ .
- 3)  $h(X) \geq g(X)$ .
- 4)  $h(X, \alpha A, Q) \geq h(X, A, Q)$  when  $\alpha > 1$ .
- 5)  $g(X, \alpha A, Q, R) \geq g(X, A, Q, R)$  when  $\alpha > 1$ .

*Proof.* 1) The first proposition is trivial.

- 2) It is easy to see that  $\tilde{g}$  is an increasing function of  $R$  and hence that  $g$  is also increasing of  $R$ . To prove that  $g$  is an increasing function of  $X$ , let us first use the matrix inversion lemma to write  $\tilde{g}$  as

$$\tilde{g}(X, R) = (X^{-1} + C^H R^{-1} C)^{-1}.$$

It is easy to see that  $\tilde{g}$  is increasing in  $X$ . Combining this with the fact that  $h$  is also an increasing function of  $X$  and  $Q$ , we conclude that  $g$  is increasing in  $X$  and  $Q$ .

- 3)

$$h(X) - g(X) = h(X)C^H(C^H h(X)C^H + R)^{-1}C^H h(X) \geq 0.$$

- 4) It is easy to see that

$$h(X, \alpha A, Q) - h(X, A, Q) = (\alpha^2 - 1)AXA^H \geq 0.$$

- 5) The proposition is true because of 4) and the fact that  $g$  is increasing in  $X$ . □

We now want to use the above propositions to simplify the problem. First, since  $g, h$  are increasing in  $Q, R$  and  $X$ , we can find  $\alpha_1 > 0$  such that  $Q \leq \alpha_1 I_n$ ,  $\Sigma \leq \alpha_1 I_n$  and  $R \leq \alpha_1 I_m$ . Since we are only concerned with the upper bound, without loss of generality we can replace  $Q, \Sigma$  and  $R$  by  $\alpha_1 I_n, \alpha_1 I_n$  and  $\alpha_1 I_m$ .

Furthermore, without loss of generality, we can assume that there is no eigenvalue of  $A$  on the unit circle by replacing  $A$  by  $(1 + \delta)A$ , where  $\delta > 0$  can be arbitrary small due to the fact that

$$h(X, \alpha A, Q) \geq h(X, A, Q), \quad g(X, \alpha A, Q, R) \geq g(X, A, Q, R),$$

if  $\alpha > 1$ .

Combining this with the assumption that  $A$  can be diagonalized, we assume that  $A = \text{diag}(A_1, A_2)$ , where  $A_i$  is diagonal and  $A_1$  and  $A_2$  contain the strictly unstable and stable eigenvalues of  $A$  respectively. We also assume  $C = [C_1, C_2]$ , where  $C_i$ s are of proper dimensions.

From the definition of  $h$  and  $g$ , we know that

$$P_k = h^{i_1} g h^{i_2 - 1} g \dots h^{i_l - 1} g h^{i_{l+1} - 1}(\Sigma),$$

provided that  $l$  packets arrive at time  $k_1 > \dots > k_l$ , where  $i_j = k_{j-1} - k_j$  for  $1 \leq j \leq l + 1$  with  $k_0 = k$ ,  $k_{l+1} = 0$ . As a result, proving Theorem 4 is equivalent to showing that

$$\text{trace}(P_k) \leq \bar{\alpha} \prod_{j=1}^l (|\lambda_j| + \varepsilon)^{2i_j}.$$

To prove the above inequality, we exploit the optimality of Kalman filtering. We construct a linear filter whose error

<sup>5</sup>By increasing we mean if  $X_1 \geq X_2$ , i.e.  $X_1 - X_2$  is positive semidefinite, then  $h(X_1) - h(X_2)$  is also positive semidefinite.

covariance satisfies (12). Therefore  $P_k$  must also satisfy (12) due to the optimality of Kalman filtering. To construct such an estimator, let us rewrite the system equations as

$$\begin{aligned} x_{k+1,1} &= A_1 x_{k,1} + w_{k,1} \\ x_{k+1,2} &= A_2 x_{k,2} + w_{k,2}, \\ y_k &= C_1 x_{k,1} + v_k + C_2 x_{k,2} \end{aligned}$$

Since  $A_2$  is stable, we can just use  $\hat{x}_{k,2} = A_2^k \bar{x}_{0,2}$  as an unbiased estimate of  $x_{k,2}$ . The new system equations become

$$\begin{aligned} x_{k+1,1} &= A_1 x_{k,1} + w_{k,1} \\ y'_k &= (y_k - C_2 A_2^k \bar{x}_{0,2}) = C_1 x_{k,1} + v_k + C_2 (x_{k,2} - A_2^k \bar{x}_{0,2}) \end{aligned}$$

To obtain an estimator for  $x_{k,1}$ , let us first write down the relation between  $y'_k$  and  $x_{k,1}$  as:

$$\begin{aligned} \begin{bmatrix} \gamma_k y'_k \\ \vdots \\ \gamma_1 y'_1 \\ \bar{x}_{0,1} \end{bmatrix} &= \begin{bmatrix} \gamma_k C_1 \\ \vdots \\ \gamma_1 C_1 A_1^{-k+1} \\ A_1^{-k} \end{bmatrix} x_{k,1} + \begin{bmatrix} \gamma_k v_k \\ \vdots \\ \gamma_1 v_1 \\ \bar{x}_{0,1} - x_{0,1} \end{bmatrix} \\ &+ \begin{bmatrix} \gamma_k C_2 (x_{k,2} - A_2^k \bar{x}_{0,2}) \\ \vdots \\ \gamma_1 C_2 (x_{1,2} - A_2 \bar{x}_{0,2}) \\ 0 \end{bmatrix} \\ &- \begin{bmatrix} 0 & \cdots & 0 & 0 \\ \gamma_{k-1} C_1 A_1^{-1} & \cdots & 0 & 0 \\ \vdots & \ddots & \vdots & \vdots \\ \gamma_1 C_1 A_1^{-k+1} & \cdots & \gamma_1 C_1 A_1^{-1} & 0 \\ A_1^{-k} & \cdots & A_1^{-2} & A_1^{-1} \end{bmatrix} \begin{bmatrix} w_{k-1,1} \\ \vdots \\ w_{0,1} \end{bmatrix}. \end{aligned} \quad (23)$$

where  $\gamma_i = 1$  if and only if  $i = k_j, j = 1, \dots, l$ . To write (23) in a more compact form, let us define the following quantities:

$$F_k \triangleq \begin{bmatrix} A_1^{-1} & \cdots & 0 & 0 \\ \vdots & \ddots & \vdots & \vdots \\ A_1^{-k+1} & \cdots & A_1^{-1} & 0 \\ A_1^{-k} & \cdots & A_1^{-2} & A_1^{-1} \end{bmatrix} \in \mathbb{C}^{lk \times lk}. \quad (24)$$

$$G_k \triangleq \begin{bmatrix} 0 & \cdots & 0 & 0 \\ C_1 & \cdots & 0 & 0 \\ \vdots & \ddots & \vdots & 0 \\ 0 & \cdots & C_1 & 0 \\ 0 & \cdots & 0 & I_l \end{bmatrix} \in \mathbb{C}^{(mk+l) \times lk}. \quad (25)$$

$$\begin{aligned} e_k &\triangleq -G_k F_k \begin{bmatrix} w_{k-1,1} \\ \vdots \\ w_{0,1} \end{bmatrix} + \begin{bmatrix} v_k \\ \vdots \\ v_1 \\ \bar{x}_{0,1} - x_{0,1} \end{bmatrix} \\ &+ \begin{bmatrix} C_2 (x_{k,2} - A_2^k \bar{x}_{0,2}) \\ \vdots \\ C_2 (x_{1,2} - A_2 \bar{x}_{0,2}) \\ 0 \end{bmatrix} \in \mathbb{C}^{mk+l}. \end{aligned} \quad (26)$$

$$T_k \triangleq \begin{bmatrix} C_1 \\ \vdots \\ C_1 A_1^{-k+1} \\ A_1^{-k} \end{bmatrix} \in \mathbb{C}^{(mk+l) \times l}, Y'_k = \begin{bmatrix} y'_k \\ \vdots \\ y'_1 \\ \bar{x}_{0,1} \end{bmatrix} \in \mathbb{C}^{mk+l}. \quad (27)$$

Define  $\Gamma_k$  as the matrix obtained by removing the zero rows from  $\text{diag}(\gamma_k I_m, \gamma_{k-1} I_m, \dots, \gamma_1 I_m, I_l)$ . Thus  $\Gamma_k$  is a  $ml+l$  by  $mk+l$  matrix. Also define

$$\tilde{Y}'_k \triangleq \Gamma_k Y'_k, \tilde{T}_k \triangleq \Gamma_k T_k, \tilde{e}_k \triangleq \Gamma_k e_k.$$

Now we can rewrite (23) in a more compact form as

$$\tilde{Y}'_k = \tilde{T}_k x_{k,1} + \tilde{e}_k. \quad (28)$$

Since we assumed, without loss of generality, that  $R, Q, \Sigma_0$  are all diagonal matrices, it is easy to see that  $x_{k,2}$  and  $x_{k,1}$  are mutually independent. As a result, one can easily prove that the following estimator of  $x_k$  is unbiased

$$\begin{aligned} \hat{x}_{k,1} &= (\tilde{T}_k^H \text{Cov}(\tilde{e}_k)^{-1} \tilde{T}_k)^{-1} \tilde{T}_k^H \text{Cov}(\tilde{e}_k)^{-1} \tilde{Y}'_k, \\ \hat{x}_{k,2} &= A_2^k \bar{x}_{0,2}, \end{aligned} \quad (29)$$

with covariance

$$P'_k = \begin{bmatrix} P'_{k,1} & P'_{k,off} \\ (P'_{k,off})^H & P'_{k,2} \end{bmatrix}, \quad (30)$$

where

$$\begin{aligned} P'_{k,1} &= (\tilde{T}_k^H \text{Cov}(\tilde{e}_k)^{-1} \tilde{T}_k)^{-1}, \\ P'_{k,2} &= \text{Cov}(x_{k,2}) = A_2^k \Sigma_0 A_2^{kH} + \sum_{i=0}^{k-1} A_2^i Q A_2^{iH}. \end{aligned} \quad (31)$$

Since we are only concerned with the trace of  $P'_k$ , it is easy to see that

$$\text{trace}(P'_k) = \text{trace}(P'_{k,1}) + \text{trace}(P'_{k,2}),$$

where  $\text{trace}(P'_{k,2})$  is uniformly bounded regardless of  $k$ . Therefore, in order to prove (12) we focus exclusively on  $\text{trace}(P'_{k,1})$ . Combining this argument with the optimality of Kalman filtering, we only need to prove that

$$\begin{aligned} \text{trace}(P'_{k,1}) &= \text{trace}[(\tilde{T}_k^H \text{Cov}(\tilde{e}_k)^{-1} \tilde{T}_k)^{-1}] \\ &\leq \bar{\alpha} \prod_{j=1}^n (|\lambda_j| + \varepsilon)^{2i_j}. \end{aligned}$$

Before we establish the result of the theorem we need to prove the following lemmas:

**Lemma 1.** If a system satisfies assumptions (H1) – (H3), then  $P'_{k,1}$  is bounded by

$$P'_{k,1} \leq \alpha_2 \left( \sum_{j=1}^l (A^{k_j-k})^H C^H C A^{k_j-k} + A^{-kH} (A^{-k})^{-1} \right), \quad (32)$$

where  $\alpha_2 \in \mathbb{R}$  is constant independent of  $i_j$ .

*Proof.* We bound  $\text{Cov}(\tilde{e}_k)$  by a diagonal matrix. Since we assume that  $w_k, v_k, x_0$  are mutually independent, it is easy to

prove that

$$\begin{aligned} \text{Cov}(e_k) &= G_k F_k \text{Cov} \left( \begin{bmatrix} w_{k-1,1} \\ \vdots \\ w_{0,1} \end{bmatrix} \right) F_k^H G_k^H \\ &+ \text{Cov} \left( \begin{bmatrix} v_k \\ \vdots \\ v_1 \\ \bar{x}_{0,1} - x_{0,1} \end{bmatrix} \right) + \text{Cov} \left( \begin{bmatrix} C_2 x_{k,2} \\ \vdots \\ C_2 x_{1,2} \\ 0 \end{bmatrix} \right). \end{aligned}$$

Let us consider the first term. Notice that

$$F_k^{-1} = \begin{bmatrix} A_1 & & & & \\ -I_l & \ddots & & & \\ & & \ddots & & \\ & & & A_1 & \\ & & & -I_l & A_1 \end{bmatrix}.$$

Therefore,

$$(F_k F_k^H)^{-1} = \begin{bmatrix} A_1^H A_1 + I & -A_1 & & & \\ -A_1^H & \ddots & \ddots & & \\ & \ddots & \ddots & A_1^H A_1 + I & -A_1 \\ & & & -A_1^H & A_1^H A_1 \end{bmatrix}.$$

By Gershgorin's Circle Theorem [21], we know that all the eigenvalues of  $(F_k F_k^H)^{-1}$  are located inside the union of the following circles:  $|\zeta - |\lambda_i|^2 - 1| = |\lambda_i|$ ,  $|\zeta - |\lambda_i|^2 - 1| = 2|\lambda_i|$ ,  $|\zeta - |\lambda_i|^2| = |\lambda_i|$ .

Since  $|\lambda_1| \geq |\lambda_2| \geq \dots \geq |\lambda_l| > 1$ , for each eigenvalue  $\zeta_i$  of  $(F_k F_k^H)^{-1}$ , the following holds<sup>6</sup>:

$$\zeta_i \geq \min\{|\lambda_i|^2 + 1 - |\lambda_i|, |\lambda_i|^2 + 1 - 2|\lambda_i|, |\lambda_i|^2 - |\lambda_i|\}. \quad (33)$$

Thus,  $0 < (|\lambda_l| - 1)^2 \leq \zeta_i$ , which in turn gives

$$F_k F_k^H \leq \frac{1}{(|\lambda_l| - 1)^2} I_{lk}.$$

Since we assume that  $Q = \alpha_1 I_n$ , we can prove that

$$\begin{aligned} G_k F_k \text{Cov} \left( \begin{bmatrix} w_{k-1,1} \\ \vdots \\ w_{0,1} \end{bmatrix} \right) F_k^H G_k^H &= \alpha_1 G_k F_k F_k^H G_k^H \\ &\leq \frac{\alpha_1}{(|\lambda_l| - 1)^2} G_k G_k^H. \end{aligned}$$

From the definition,

$$\frac{\alpha_1}{(|\lambda_l| - 1)^2} G_k G_k^H = \frac{\alpha_1}{(|\lambda_l| - 1)^2} \text{diag}(0, C_1 C_1^H, \dots, C_1 C_1^H, I_l), \quad i \neq a.$$

which is uniformly bounded by  $\alpha_3 I_{mk+l}$ .

Let us now consider the second term. Since  $R = \alpha_1 I_m$ , it is trivial to see that

$$\text{Cov} \left( \begin{bmatrix} v_k \\ \vdots \\ v_1 \\ \bar{x}_{0,1} - x_{0,1} \end{bmatrix} \right) = \alpha_1 I_{mk+l}.$$

<sup>6</sup>Note that all the eigenvalues of  $(F_k F_k^H)^{-1}$  are real.

Consider now the last term. Write  $x_{k,2}$  as

$$\begin{bmatrix} x_{k,2} \\ \vdots \\ x_{0,2} \end{bmatrix} = \begin{bmatrix} I_{n-l} & \cdots & A_2^{k-1} & A_2^k \\ \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & I_{n-l} & A_2 \\ 0 & \cdots & 0 & I_{n-l} \end{bmatrix} \begin{bmatrix} w_{k-1,2} \\ \vdots \\ w_{0,2} \\ x_{0,2} \end{bmatrix}.$$

As a result,

$$\begin{aligned} \text{Cov} \left( \begin{bmatrix} x_{k,2} \\ \vdots \\ x_{0,2} \end{bmatrix} \right) &= \alpha_1 \begin{bmatrix} I_{n-l} & \cdots & A_2^k \\ \vdots & \ddots & \vdots \\ 0 & \cdots & I_{n-l} \end{bmatrix} \begin{bmatrix} I_{n-l} & \cdots & A_2^k \\ \vdots & \ddots & \vdots \\ 0 & \cdots & I_{n-l} \end{bmatrix}^H \\ &\leq \frac{\alpha_1}{(1 - |\lambda_{l+1}|)^2} I_{(n-l)(k+1)}, \end{aligned}$$

where the proof of the last inequality is similar to the proof of  $F_k F_k^H \leq (|\lambda_l| - 1)^{-2} I_{lk}$  and is omitted. Therefore it is easy to see that

$$\text{Cov} \left( \begin{bmatrix} C_2 x_{k,2} \\ \vdots \\ C_2 x_{1,2} \\ 0 \end{bmatrix} \right) \leq \alpha_4 I_{mk+l},$$

where  $\alpha_4$  is a constant. As a result we have proved that

$$\text{Cov}(e_k) \leq \alpha_2 I_{mk+l},$$

where  $\alpha_2 = \alpha_3 + \alpha_1 + \alpha_4$ . Moreover

$$\text{Cov}(\tilde{e}_k) = \Gamma_k \text{Cov}(e_k) \Gamma_k^H \leq \alpha_2 I_{ml+l}.$$

The above bound is independent of  $i_1, \dots, i_{l+1}$ , which proves

$$\begin{aligned} P'_{k,1} &= (\tilde{T}_k^H \text{Cov}(\tilde{e}_k)^{-1} \tilde{T}_k)^{-1} \\ &\leq \alpha_2 \left( \sum_{j=1}^l (A^{k_j-k})^H C^H C A^{k_j-k} + A^{-kH} A^{-k} \right)^{-1}. \end{aligned}$$

□

We manipulate  $\sum_{j=1}^n (A^{-i_j})^H C^H C A^{-i_j}$  to prove the upper bound by using cofactors for matrix inversion. Before continuing, we need the following lemmas.

**Lemma 2.** For a non-degenerate system, it is possible to find a set of row vectors  $L_1, L_2, \dots, L_l$ , such that  $L_i C = [l_{i,0}, \dots, l_{i,l}]$ , where  $l_{i,i} = 1$  and  $l_{i,a} = 0$  if  $|\lambda_i| = |\lambda_a|$  and  $i \neq a$ .

*Proof:* It is simple to show that the lemma holds by using Gaussian Elimination for every quasi-equiblock. ■

**Lemma 3.** Consider that  $|\lambda_1| \geq |\lambda_2| \geq \dots \geq |\lambda_l|$ ,  $l_{i,i} = 1$  and  $l_{i,a} = 0$  if  $i \neq a$  and  $|\lambda_i| = |\lambda_a|$ . Let  $i_1 = k - k_1$  and  $i_j = k_{j-1} - k_j$  for  $2 \leq j \leq l$ . Define

$$D = \begin{bmatrix} l_{1,1} \lambda_1^{k_1-k} & l_{1,2} \lambda_2^{k_1-k} & \cdots & l_{1,l} \lambda_l^{k_1-k} \\ l_{2,1} \lambda_1^{k_2-k} & l_{2,2} \lambda_2^{k_2-k} & \cdots & l_{2,l} \lambda_l^{k_2-k} \\ \vdots & \vdots & \ddots & \vdots \\ l_{l,1} \lambda_1^{k_l-k} & l_{l,2} \lambda_2^{k_l-k} & \cdots & l_{l,l} \lambda_l^{k_l-k} \end{bmatrix}.$$

Then  $D$  is asymptotic to  $\prod_{j=1}^l \lambda_j^{k_j-k}$ , i.e.

$$\lim_{i_1, i_2, \dots, i_l \rightarrow +\infty} \frac{D}{\prod_{j=1}^l \lambda_j^{k_j-k}} = 1. \quad (34)$$

*Proof of Lemma 3:* The determinant  $D$  has  $l!$  terms, which have the form  $\text{sgn}(\sigma) \prod_{j=1}^l l_{j,a_j} \lambda_{a_j}^{k_j-k}$ .  $\sigma = (a_1, a_2, \dots, a_l)$  is a permutation of the set  $\{1, 2, \dots, l\}$  and  $\text{sgn}(\sigma) = \pm 1$  is the sign of permutation. Rewrite (34) as

$$\begin{aligned} \frac{D}{\prod_{j=1}^l \lambda_j^{k_j-k}} &= \sum_{\sigma} \text{sgn}(\sigma) \frac{\prod_{j=1}^l l_{j,a_j} \lambda_{a_j}^{k_j-k}}{\prod_{j=1}^l \lambda_j^{k_j-k}} \\ &= \sum_{\sigma} \text{sgn}(\sigma) \prod_{j=1}^l l_{j,a_j} \frac{\left(\prod_{j=1}^l \lambda_{a_j}\right)^{-i_1} \cdots \left(\prod_{j=l}^l \lambda_{a_j}\right)^{-i_l}}{\left(\prod_{j=1}^l \lambda_j\right)^{-i_1} \cdots \left(\prod_{j=l}^l \lambda_j\right)^{-i_l}} \\ &= \sum_{\sigma} \text{sgn}(\sigma) \prod_{j=1}^l l_{j,a_j} \prod_{b=1}^l \left(\frac{\prod_{j=b}^l \lambda_{a_j}}{\prod_{j=b}^l \lambda_j}\right)^{-i_j}. \end{aligned}$$

Now we can just analyze each term of the sum. Since  $|\lambda_1| \geq \dots \geq |\lambda_l|$ , then  $|\prod_{j=b}^l \lambda_{a_j}| \geq |\prod_{j=b}^l \lambda_j|$ . First consider that there exist some  $j$ s such that  $|\lambda_{a_j}| \neq |\lambda_j|$  and define  $j^*$  to be the largest, which means  $|\lambda_{a_{j^*}}| \neq |\lambda_{j^*}|$  and  $|\lambda_{a_j}| = |\lambda_j|$  for all  $j > j^*$ . Since  $|\lambda_{j^*}|$  is the smallest among  $|\lambda_1|, \dots, |\lambda_j|$ , we know that  $|\lambda_{a_{j^*}}| > |\lambda_{j^*}|$ . Thus,

$$\begin{aligned} \left| \frac{\prod_{j=j^*}^l \lambda_{a_j}}{\prod_{j=j^*}^l \lambda_j} \right| &> 1, \\ \lim_{i_1, i_2, \dots, i_l \rightarrow \infty} \left| \prod_{j=1}^l l_{j,a_j} \prod_{b=1}^l \left(\frac{\prod_{j=b}^l \lambda_{a_j}}{\prod_{j=b}^l \lambda_j}\right)^{-i_j} \right| \\ &\leq \left| \prod_{j=1}^l l_{j,a_j} \right| \lim_{i_{j^*} \rightarrow \infty} \left| \frac{\prod_{j=j^*}^l \lambda_{a_j}}{\prod_{j=j^*}^l \lambda_j} \right|^{-i_{j^*}} = 0. \end{aligned}$$

Then consider that if for all  $j$ ,  $|\lambda_{a_j}| = |\lambda_j|$ , but  $(a_1, \dots, a_l) \neq (1, 2, \dots, l)$ . Thus, there exists  $j^*$  such that  $a_{j^*} \neq j^*$ . Hence  $l_{j^*, a_{j^*}} = 0$ . Therefore, these terms are always 0.

The only term left is

$$\text{sgn}(\sigma) \prod_{j=1}^l l_{j,j} \prod_{b=1}^l \left(\frac{\prod_{j=b}^l \lambda_j}{\prod_{j=b}^l \lambda_j}\right)^{-i_j} = 1.$$

Thus, we can conclude that

$$\lim_{i_1, i_2, \dots, i_l \rightarrow \infty} \frac{D}{\prod_{j=1}^l \lambda_j^{k_j-k}} = 1.$$

Since the system is non-degenerate, from Lemma 2, we know that there exist  $L_1, L_2, \dots, L_l$ , such that  $L_i C = [l_{i,1}, \dots, l_{i,l}]$  is a row vector,  $l_{i,i} = 1$  and  $l_{i,a} = 0$  if  $i \neq a$  and  $|\lambda_i| = |\lambda_a|$ .

Define matrices

$$U \triangleq \begin{bmatrix} l_{1,1} \lambda_1^{-i_1} & l_{1,2} \lambda_2^{-i_1} & \cdots & l_{1,l} \lambda_l^{-i_1} \\ l_{2,1} \lambda_1^{-i_2} & l_{2,2} \lambda_2^{-i_2} & \cdots & l_{2,l} \lambda_l^{-i_2} \\ \vdots & \vdots & \ddots & \vdots \\ l_{l,1} \lambda_1^{-i_l} & l_{l,2} \lambda_2^{-i_l} & \cdots & l_{l,l} \lambda_l^{-i_l} \end{bmatrix}, O \triangleq U^{-1}. \quad (35)$$

Define  $\alpha_5 = \max(\lambda_{\max}(L_1^H L_1), \dots, \lambda_{\max}(L_l^H L_l))$ . Thus,  $L_i^H L_i \leq \alpha_5 I_m$ , and

$$\begin{aligned} \sum_{j=1}^l A^{-i_j H} C^H C A^{-i_j} &\geq \sum_{j=1}^l \frac{1}{\alpha_5} A^{-i_j H} C^H L_j^H L_j C A^{-i_j} \\ &= \frac{1}{\alpha_5} \begin{bmatrix} A^{-i_1 H} C^H L_1^H & \cdots & A^{-i_l H} C^H L_l^H \end{bmatrix} \begin{bmatrix} L_1 C A^{-i_1} \\ \vdots \\ L_l C A^{-i_l} \end{bmatrix} \\ &= \frac{1}{\alpha_5} U^H U, \end{aligned} \quad (36)$$

and

$$\begin{aligned} \left( \sum_{j=1}^l A^{(k_j-k)H} C^H C A^{k_j-k} \right)^{-1} &\leq \alpha_5 (U^H U)^{-1} \\ &= \alpha_5 O O^H \leq \alpha_5 \text{trace}(O O^H) I_l = \alpha_5 \sum_{a,b} O_{a,b} (O^H)_{b,a} I_l \\ &= \alpha_5 \sum_{a,b} O_{a,b} \times \text{conj}(O_{a,b}) I_l = \alpha_5 \sum_{a,b} |O_{a,b}|^2 I_l, \end{aligned} \quad (37)$$

where  $\text{conj}()$  means complex conjugate.

From Lemma 3, we can compute the cofactor matrix of  $U$  and hence  $O = U^{-1}$ . Define the minor  $M_{a,b}$  of  $U$  as the  $(l-1) \times (l-1)$  matrix that results from deleting row  $a$  and column  $b$ . Thus

$$O_{a,b} = \frac{(-1)^{a+b} \det(M_{b,a})}{\det(U)}. \quad (38)$$

From Lemma 3 we know that

$$\lim_{i_1, i_2, \dots, i_l \rightarrow \infty} \frac{\det(U)}{\prod_{j=1}^l \lambda_j^{k_j-k}} = 1.$$

Since  $M_{a,b}$  has the same structure as  $U$ , it is easy to show that

$$\det(M_{a,b}) \leq \rho_{a,b} \prod_{j=2}^l |\lambda_j^{k_j-1-k}|,$$

where  $\rho_{a,b}$  is a constant. Thus,

$$\begin{aligned}
& \limsup_{i_1, \dots, i_l \rightarrow \infty} \frac{\left( \sum_{j=1}^l A^{k_j - kH} C^H C A^{k_j - k} \right)^{-1}}{\prod_{j=1}^l |\lambda_j|^{2i_j}} \\
& \leq \limsup_{i_1, \dots, i_l \rightarrow \infty} \frac{\alpha_5 \sum_{a,b} |O_{a,b}|^2}{\prod_{j=1}^l |\lambda_j|^{2i_j}} I_l \\
& = \limsup_{i_1, \dots, i_l \rightarrow \infty} \alpha_5 \left( \sum_{a,b} \left| \frac{\det(M_{a,b})}{\det(U)} \right|^2 / \prod_{j=1}^l |\lambda_j|^{2i_j} \right) I_l \quad (39) \\
& \leq \alpha_5 \left( \sum_{a,b} \rho_{a,b}^2 \left| \frac{\prod_{j=2}^l |\lambda_j^{k_j - 1 - k}|}{\prod_{j=1}^l \lambda_j^{k_j - k}} \right|^2 / \prod_{j=1}^l |\lambda_j|^{2i_j} \right) I_l \\
& = \alpha_5 \sum_{i,j} \rho_{i,j}^2 I_l.
\end{aligned}$$

Hence, there exists  $\alpha_6 > 0$  such that

$$\left( \sum_{j=1}^l A^{(k_j - k)H} C^H C A^{k_j - k} \right)^{-1} \leq \alpha_6 \prod_{j=1}^l |\lambda_j|^{2i_j} I_l,$$

Along with Lemma 1 we conclude the proof of Theorem 4.

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